

**ECE 730, Lec. 1**  
**Exam 1**  
**Wednesday, 17 Oct. 2018**  
**5:00 pm – 6:30 pm**  
**2534 EH**

**100 Points**

**Justify your answers!**

**Be precise!**

**Closed Book**

**Closed Notes**

**No Calculators**

**You may bring one sheet of  $8.5 \times 11$  paper with notes written on both sides.**

The solutions of a quadratic equation  $as^2 + bs + c = 0$  are

$$s = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}.$$

1. [20 pts.] Let  $X$  and  $Y$  be random vectors, and suppose you are given

$$\begin{aligned} m_X &:= E[X] & R_X &:= E[XX'] & \text{and} & R_{XY} &:= E[XY'] \\ m_Y &:= E[Y] & R_Y &:= E[YY'] \end{aligned}$$

Express the **linear** MMSE estimate of  $X$  based on  $Y$  using any of the above quantities as appropriate. If you need the inverse of a matrix, assume it exists.

**Solution.** We know that the required estimate is  $A(Y - m_Y) + m_X$ , where  $A$  solves  $AC_Y = C_{XY}$ . Since

$$C_Y = R_Y - m_Y m_Y', \quad \text{and} \quad C_{XY} = R_{XY} - m_X m_Y',$$

we have

$$A = (R_{XY} - m_X m_Y')(R_Y - m_Y m_Y')^{-1}.$$

So the estimate is

$$(R_{XY} - m_X m_Y')(R_Y - m_Y m_Y')^{-1}(Y - m_Y) + m_X.$$

2. [20 pts.] Let  $X \sim \exp(\lambda)$ , and given  $X = x$ , let the conditional density of  $Y$  be  $N(x^3, \sigma^2)$ . Compute  $E[Y^2]$ . **Show your work.**

**Solution.**  $E[Y^2] = E[E[Y^2|X]] = E[\sigma^2 + (X^3)^2] = \sigma^2 + 6/\lambda^6.$

3. [20 pts.] Let  $Y = (X + W)^2$ , where  $X \sim N(0, 0.9)$ ,  $W \sim N(0, 0.1)$ , and  $X$  and  $W$  are independent. Find the density of  $Y$ . **Justify your answer.**

**Solution.** Since  $X$  and  $W$  are independent Gaussians,  $X + W$  is also Gaussian, with zero mean and variance  $0.9 + 0.1 = 1$ , by Problem 4.55(a). Hence,  $Y$  is chi-squared with one degree of freedom by Problem 4.46.

4. [20 pts.] If  $X \sim \exp(\lambda)$  and  $Y \sim \exp(\mu)$  are independent, find the density of  $Z := \max(X, Y)$ . **Show your work.**

**Solution.** Since  $Z \geq 0$ , it suffices to consider  $z > 0$ . For such  $z$ ,

$$F_Z(z) = P(Z \leq z) = P(\max(X, Y) \leq z) = P(X \leq z, Y \leq z) = F_X(z)F_Y(z).$$

Then

$$\begin{aligned} f_Z(z) &= \frac{d}{dz} F_Z(z) = f_X(z)F_Y(z) + F_X(z)f_Y(z) = \lambda e^{-\lambda z}(1 - e^{-\mu z}) + (1 - e^{-\lambda z})\mu e^{-\mu z} \\ &= \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda + \mu)z}. \end{aligned}$$

5. [20 pts.] Suppose  $X$  is a two-dimensional random vector with zero mean and covariance matrix

$$C_X = \begin{bmatrix} \varepsilon & 1 \\ 1 & \varepsilon \end{bmatrix},$$

where  $\varepsilon > 1$  is a parameter. Let  $X = PY$  be the Karhunen–Loève expansion of  $X$ . Find the covariance matrix of  $Y$ . Express your answer in terms of  $\varepsilon$ . *Hint:* This does not require finding the transformation  $P$ .

**Solution.** Since the covariance matrix of  $Y$  is diagonal with diagonal elements being the eigenvalues of  $C_X$ , it suffices to solve the characteristic equation

$$\det(sI - C_X) = \det \begin{bmatrix} s - \varepsilon & -1 \\ -1 & s - \varepsilon \end{bmatrix} = s^2 - 2\varepsilon s + \varepsilon^2 - 1 = 0. \quad (*)$$

By the quadratic formula,

$$s = \frac{2\varepsilon \pm \sqrt{4\varepsilon^2 - 4(\varepsilon^2 - 1)}}{2} = \varepsilon \pm 1,$$

and follows that

$$C_Y = \begin{bmatrix} \varepsilon + 1 & 0 \\ 0 & \varepsilon - 1 \end{bmatrix}. \quad (**)$$

Rather than using the quadratic formula, we could have rearranged the last equation in (\*) as  $(s - \varepsilon)^2 = 1$ . Taking square roots yields  $|s - \varepsilon| = 1$ , or  $s - \varepsilon = \pm 1$ , which says that  $s = \varepsilon \pm 1$ .

**Alternative Solution.** The eigenvalues of  $C_X$  solve

$$\begin{bmatrix} \varepsilon & 1 \\ 1 & \varepsilon \end{bmatrix} \begin{bmatrix} u \\ v \end{bmatrix} = \lambda \begin{bmatrix} u \\ v \end{bmatrix},$$

where  $[u, v]^T$  is not the zero vector. From the second equation,  $u = (\lambda - \varepsilon)v$ . Note that this implies  $v \neq 0$ , since otherwise  $[u, v]^T$  is the zero vector. Now rewrite the first equation as

$$\begin{aligned} v &= (\lambda - \varepsilon)u = (\lambda - \varepsilon)(\lambda - \varepsilon)v \\ &= (\lambda - \varepsilon)^2 v. \end{aligned}$$

Rewriting this as  $[(\lambda - \varepsilon)^2 - 1]v = 0$ , and recalling that  $v \neq 0$ , we must have  $(\lambda - \varepsilon)^2 = 1$  as in the above solution. So we again have that  $C_Y$  is given by (\*\*).